

# Order Matters (?): Alternatives to Conventional Practices for Ordinal Categorical Response Variables

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## ***Ordinal Response Variables***

- Ordinal scales are prevalent in survey research
- Frequently used as response variables
- Interest in  $Y = f(\mathbf{x}'\beta)$
- Normal-theory methods commonly applied (OLS)
- Cumulative link models also applied
- Either strategy often works poorly

## ***Common Issues That Arise***

- Equal-interval scoring usually unrealistic
- “Parallel” regression assumption frequently will not hold
- Ordinality may not exist, conditional on covariates  $x$
- Alternative contrasts may be of interest
- Several alternative models are considered

## Utility Motivation

- Assume  $Y$  is a discretized measure of  $Y^*$
- Postulate  $Y_i^* = \alpha + \mathbf{x}'\beta + \epsilon_i$
- Cut Point Rule:

$$Y = 1 \quad \equiv \quad Y^* \leq \alpha_1$$

$$Y = 2 \quad \equiv \quad \alpha_1 < Y^* \leq \alpha_2$$

$$Y = 3 \quad \equiv \quad \alpha_2 < Y^* \leq \alpha_3$$

$$Y = 4 \quad \equiv \quad Y^* > \alpha_3$$

- Specify C.D.F. for  $\epsilon$  (logistic, standard normal, cloglog)
- With logistic, proportional odds is obtained.

## Proportional Odds Model

- Gives rise to:

$$\Pr(Y \leq y_j \mid \mathbf{x}) = \frac{\exp(\alpha_j - \mathbf{x}'\beta)}{1 + \exp(\alpha_j - \mathbf{x}'\beta)}$$

- Linear Model for log-odds:

$$\log \left[ \frac{\Pr(Y \leq y_j \mid \mathbf{x})}{\Pr(Y > y_j \mid \mathbf{x})} \right] = \alpha_j - \mathbf{x}'\beta, \quad j = 1, 2, \dots, j-1$$

- Proportional Odds Property:

$$\frac{\exp(x_1\beta)}{\exp(x_2\beta)} = \exp\{(x_1 - x_2)'\beta\}$$

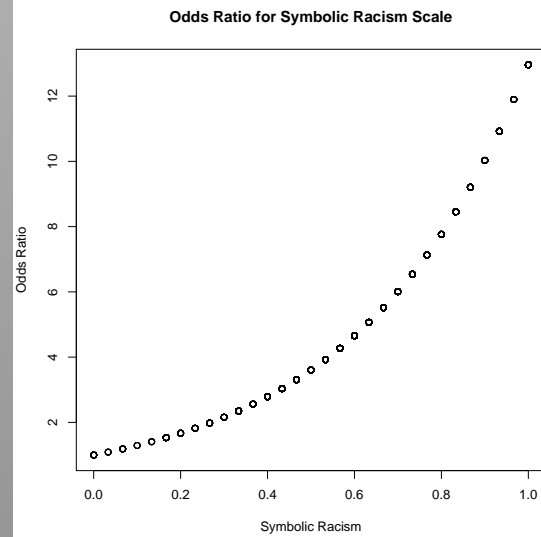
## ***Standard Practice***

- Assume  $Y$  is equal-interval scored and use normal-theory methods (OLS)
- Apply cumulative link model but rarely evaluate prop. odds assumption
- Or do both, conduct “eyeball test”, and report OLS

**Table 1: Support for Affirmative Action  
OLS and Proportional Odds Estimates**

Variable	Estimate	
	OLS	Proportional Odds
Symbolic Racism	1.375 (0.111)	2.562 (0.207)
Racial Prejudice	0.383 (0.193)	0.728 (0.356)
Education	0.259 (0.102)	0.510 (0.182)
Ideology	0.066 (0.027)	0.117 (0.047)
Female	0.025 (0.049)	0.075 (0.088)
Intercept	1.479 (0.127)	
$\alpha_1$		0.652 (0.233)
$\alpha_2$		1.888 (0.236)
$\alpha_3$		3.268 (0.245)
$n$ log-likelihood:	1744	1744 -2289.518

# Symbolic Racism



- Odds:  $\exp(2.562) = 12.96$
- Odds Ratios Invariant to the Cut Points
- and they are proportional
- ... but does the property hold?

## Proportionality Tests

**Table 2: Testing the Proportional Odds Assumption**

Variable	Coefficient		
	$y > 1$	$y > 2$	$y > 3$
Symbolic Racism	1.822	2.458	3.037
Racial Prejudice	1.384	0.219	0.664
Education	0.209	0.496	0.651
Ideology	0.116	0.105	0.139
Female	-0.118	0.144	0.052
Constant	-0.279	-1.651	-3.645

Variable	$\chi^2$	$p > \chi^2$	df
All	29.33	0.001	10
Symbolic Racism	12.37	0.002	2
Racial Prejudice	8.15	0.017	2
Education	2.29	0.319	2
Ideology	0.33	0.849	2
Female	5.79	0.055	2

## Nonproportional Odds

- Generalized Model:  $\log \left[ \frac{\Pr(Y \leq y_j | \mathbf{x})}{\Pr(Y > y_j | \mathbf{x})} \right] = \alpha_j - \mathbf{x}'\beta_j, \quad j = 1, 2, \dots, j - 1$

- Partial Proportional Odds:

$$\log \left[ \frac{\Pr(Y \leq y_j | \mathbf{x})}{\Pr(Y > y_j | \mathbf{x})} \right] = -\alpha_j - \mathbf{x}'\beta - \mathbf{t}'\gamma_j, \quad j = 1, 2, \dots, j - 1$$

- Restricted Generalized Logit:

$$\log \left[ \frac{\Pr(Y \leq y_j | \mathbf{x})}{\Pr(Y > y_j | \mathbf{x})} \right] = \alpha_j + \mathbf{x}'\beta + \mathbf{z}'\zeta_j, \quad j = 1, 2, \dots, j - 1$$

**Table 3: Support for Affirmative Action  
Partial Proportional Odds (UPP Model)**

<b>Variable</b>	<b>Coefficient</b>	<b>Standard Error</b>
Symbolic Racism	1.788	(0.285)
$\gamma_2$	0.800	(0.257)
$\gamma_3$	1.345	(0.358)
Racial Prejudice	1.563	(0.495)
$\gamma_2$	-1.437	(0.440)
$\gamma_3$	-0.811	(0.586)
Female	-0.098	(0.125)
$\gamma_2$	0.277	(0.111)
$\gamma_3$	0.147	(0.145)
Education	0.479	(0.182)
Ideology	0.122	(0.047)
$\alpha_1$	.273	(.22)
$\alpha_2$	-1.635	(.203)
$\alpha_3$	-3.263	(.236)
$n$	1744	
log-likelihood	-2274.293	

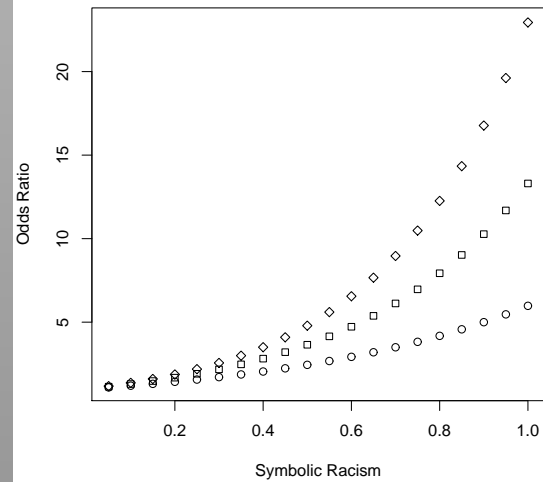
## Application: Restricted Generalized Ordinal

Table 3a: Support for Affirmative Action

Partial Proportional Odds (Generalized Ordinal Logit)

Variable	Coefficient		
	$C_1$	$C_2$	$C_3$
Symbolic Racism	1.788 (0.285)	2.588 (0.245)	3.133 (0.280)
Racial Prejudice	1.563 (0.495)	0.127 (0.421)	0.752 (0.462)
Education	0.478 (0.182)	0.478 (0.182)	0.478 (0.182)
Ideology	0.122 (0.047)	0.122 (0.047)	0.122 (0.047)
Female	-0.098 (0.125)	0.180 (0.101)	0.049 (0.113)
Constant	-0.509 (0.283)	-1.698 (0.261)	-3.639 (0.299)
$n$ log-likelihood	1744 -2274.293		

## Symbolic Racism Redux



- Odds vary over cut points
- “Parallel Regression” does not hold
- Fit of partial prop. odds superior to proportional odds

## *Alternative Modeling Strategies*

- Sometimes ordinality will not hold (or will not be of primary interest)
- Different contrasts may be of interest
- Leads to consideration of alternative models

## Multinomial Models

- Baseline Category Logit:  $\log \left[ \frac{\Pr(Y=y_j|\mathbf{x})}{\Pr(Y=1|\mathbf{x})} \right] = \alpha_j + \mathbf{x}'\beta_j, \quad j = 1, 2, \dots, j-1$
- Adjacent Category Logit:  $\log \left[ \frac{\Pr(Y=y_{j+1}|\mathbf{x})}{\Pr(Y=y_j|\mathbf{x})} \right] = \alpha_j + \mathbf{x}'\beta_j, \quad j = 1, 2, \dots, j-1$
- Anderson's Stereotype Logit:  $\log \left[ \frac{\Pr(Y=y_j|\mathbf{x})}{\Pr(Y=1|\mathbf{x})} \right] = \alpha_j + \mathbf{x}'\phi_j\beta, \quad j = 1, 2, \dots, j-1$
- Latter gives a test of ordinality:  $1 = \phi_1 > \phi_2 > \dots \phi_J = 0$

**Application: ACL**

**Table 4:  
Adjacent Category Logit**

<b>Variable</b>	<b>Coefficient</b>		
	2 vs 1	3 vs 2	4 vs 3
Moral Traditionalism	-0.812 (0.776)	-1.517 (0.639)	-1.752 (0.742)
Group Difference	-0.010 (0.096)	-0.028 (0.076)	-0.001 (0.088)
Hispanic Traits	-0.417 (0.948)	-1.092 (0.779)	-1.928 (0.876)
Ideology	0.490 (0.317)	0.031 (0.262)	0.027 (0.304)
Economic Evaluation	-0.324 (0.402)	-0.179 (0.336)	-0.714 (0.410)
Constant	1.161 (0.748)	2.235 (0.603)	1.353 (0.639)
n	774		
log-likelihood	-981.839		

## ***Contrasts are for Adjacent Categories***

- Consider the Moral Traditionalism Scale
- C2 vs. C1: .44
- C3 vs. C2: .22
- C4 vs. C3: .17
- Odds are decreasing over each adjacent category
- Note that  $ACL=BCL$ : No statistical adjudication is possible
- Just the interpretation differs: ordinality is preserved through local logits
- ... though ordinality *need* not hold

## Application: Stereotype

Table 5:  
Stereotype Logit Model

Variable	Coefficient
Moral Traditionalism	4.111 (0.852)
Group Difference	0.044 (0.093)
Hispanic Traits	3.444 (0.990)
Ideology	-0.410 (0.328)
Economic Evaluation	1.092 (0.437)
$\alpha_1$	-4.725 (0.825)
$\alpha_2$	-3.660 (0.726)
$\alpha_3$	-1.382 (0.624)
$\phi_1$	1
$\phi_2$	0.848 (0.131)
$\phi_3$	0.507 (0.102)
$\phi_4$	0
$n$	774
log-likelihood	-983.586

## *Interpretation?*

- Ordinality condition holds (restrictions could be applied in some instances)
- Odds Ratios are in reference to baseline category ( $Y = 4$ )
- C1 vs. C4: 61 (Constrast of Scale Extremes)
- C2 vs. C4: 33
- C3 vs. C4: 8
- Model is not equivalent to BCL (unless  $K = 1$ )
- It is reduced rank BCL
- LR Test suggests BCL does not fit better

## ***Conclusions***

- Applied work frequently makes assumptions that will not hold
- A lot of potentially interesting information is ignored
- Different models yield different contrasts
- These alternatives may be more consistent with some theories of political behavior